# \$FMSO

#### Education AFM 113 Final Prep

Disclosure: This material is for educational purposes only and is intended to supplement course content. Please ensure you review the class materials independently.

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### ŠFmso

#### 1.0 Simple Linear Regression

#### **1.1 Linear Regression Manual Process**

**Simple Linear Regression:** Analysis only includes one independent variable and the relationship between the independent (X) and dependent (Y) variables is represented by a straight line.

Exact Model:

- $Y = \alpha + \beta X + \varepsilon$ 
  - *Alpha* ( $\alpha$ ) represents the intercept.
  - *Beta* ( $\beta$ ) represents the coefficient of the independent variable (the slope).
  - *Epsilon* ( $\epsilon$ ) represents the random distances from the individual points to the best fitting line (residual terms).
- <u>Goal:</u> Identifying the best fit that will represent the relationship between X and Y in an equation as accurately as possible (measure the best line that fits the data).

Prediction Model:

• 
$$\hat{\mathbf{Y}} = \alpha + \beta \mathbf{X}$$

- *Y-hat*  $(\hat{Y})$  represents the predicted model dependent values.
- $\circ \quad Y = \hat{Y} + \epsilon$

Sum of Squared Residuals (Least Squares Method):

$$SSR = \sum_{i=1}^{N} (\alpha + \beta x_i - y_i)^2$$

Mean (µ):

$$\mu = \left(\sum_{i=1}^{N} x_{i}\right)/N$$

- -



Standard Deviation ( $\sigma$ ):

$$\sigma = \sqrt{\frac{\sum_{i=1}^{N} (x_i - \mu)^2}{N}}$$

#### **1.2 Linear Regression R Code**

summary(lm(dependent\_variablename ~ independent\_variablename, dataset)) ^ Results will reveal the intercept ( $\alpha$ ) & the slope/coefficient ( $\beta$ ).

#### 2.0 Probability Density Curve

#### 2.1 Probability Density Curve Manual Process

**Probability Density Curve:** Visualizes the probability distribution, allows us to see how probabilities are distributed over the values of a random variable.

Features of Density Curves:

- A density curve must lie on or above the horizontal axis.
- Area under the density curve (between curve & horizontal axis) must always equal 1 or 100%.
- Probability density will always be between 0 & 1 (area under the curve).
- <u>Probability density  $\neq$  probability.</u>
  - <u>Probability density</u> compares the likelihood of observing a value.
  - <u>Probability</u> compares the probability of the value falling in a range of observations.
- For a continuous variable (-∞ < x < ∞), discussing its probability of being a specific value is not meaningful because it always equals 0.

$$\circ \quad \Pr(\mathbf{x}=\#)=\mathbf{0}$$

**Outliers:** A data point that significantly deviates from the general pattern or average of the rest of the data points in a dataset.

- Mean > Median  $\Rightarrow$  Right-Skewed (upper)
- Mean < Median  $\Rightarrow$  Left-Skewed (lower)



#### 2.11 General Normal Distribution

 $X \sim N(\mu, \sigma^2)$ 

The normal distribution is characterized by its mean ( $\mu$ ) and/or standard deviation ( $\sigma$ ).

- Mean tells us where the center of the curve is.
- Standard deviation tells us how wide the curve is (will also determine the height of the curve).

#### 2.12 Standard Normal Distribution

#### Z ~ N (0, 1)

Properties of the standard normal distribution (Z probabilities):

- If Pr(-n < Z < n), then = Pr(z < n) Pr(z < -n).
- If Pr(z < 0), then = Pr(z > 0) = 50%.
- If Pr(z > a), then = Pr(z < -a).
- If Pr(z > a), then = 1-Pr(z < a).

**Standardization:** Transforming a general normal distribution into the standard normal distribution.

•  $Z = (X - \mu) / \sigma$ 

#### 2.2 Percentage Returns R Code

```
Returns <- dataset %>% mutate(dependent_returns = (dependentvariablename -
lag(dependentvariablename)) / lag(dependentvariablename), independent_returns
= (independentvariablename - lag(independentvariablename)) /
lag(independentvariablename)) %>% na.omit()
```

^ Results will give the percentage returns of the variables.

^ lag() ← Captures end-of-previous month prices.

#### 2.3 Beta Rolling Window R Code

```
rollBeta <- data.frame(WindowEndMonth = as.Date(character()), beta =
numeric(), stringsAsFactors = FALSE)
for(i in 1:(nrow(dataset) - #-1)) {
subset_df <- dataset[i:(i+#-1), ]
subset_beta <-lm(dependent_variablename ~ independent_variablename,
subset_df)
beta_coef <- coef(subset_beta)["independent_variablename"]
rollBeta <- rbind(rollBeta, data.frame(WindowEndMonth =
subset_df$datevariablename[#],
beta = beta_coef, row.names=NULL)) }</pre>
```



^ Identifies an estimated beta value for a dependent variable using a rolling window cycle (automatic rolling window on  $R \rightarrow$  line 2&3, remove line 2&3 for manual).

#### 2.4 Graph Generation R Code

```
binwidth <- (max(dataset$variablename) - min(dataset$variablename))/#ofbins
bin_edges <-</pre>
```

seq(min(dataset\$variablename),max(dataset\$variablename),binwidth)

^ Determine the cutoff values for histogram bins manually (display range of returns for each bin).

```
Ggplot(dataset, aes(x = variablename, y =..density..)) + geom_point() +
geom_smooth(method = "lm", se = False) + geom_histogram(breaks = bin_edges,
fill = "color", alpha = #) + geom_line() + geom_density(color = "color", size
= #) + geom_qq() + geom_vline(xintercept = #, linetype = "type") +
geom_abline() + scale_axis_continuous(labels = scales::percent) + labs(title
= "title", x= "dependent_variablename", y = "independent_variablename") +
theme(plot.title = element_text(hjust = #), axis.text/title.axis =
element_text(angle = #)
```

^ geom\_point()  $\leftarrow$  scatterplot, geom\_smooth()  $\leftarrow$  smooth line, geom\_histogram()  $\leftarrow$  histogram, geom\_line()  $\leftarrow$  time-series plot, geom\_density()  $\leftarrow$  probability density curve, geom\_qq()  $\leftarrow$  quantile-quantile plot, geom\_vline()  $\leftarrow$  vertical line, geom\_abline()  $\leftarrow$  reference line (intercept = 0, slope = 1).

^ scale\_axis\_continuous()  $\leftarrow$  Specifies which axis to scale continuously.

```
^ theme_classic() \leftarrow no grid lines, theme_bw() \leftarrow grey grid lines, theme() \leftarrow Customized looks.
^ se = False \leftarrow without the confidence interval, ...density... \leftarrow scale y-axis from frequency to
probability density, scales::percent \leftarrow Label the scales in percentage.
```

#### 2.3 Probability Calculation R Code

```
pnorm(x_value, mean = #, sd = #)
```

^ Results will identify the probability at x-value (z value, mean = 0, sd = 1).

 $^{\rm o}$  dnorm()  $\leftarrow$  gives probability density, qnorm()  $\leftarrow$  finds quartile/percentile.

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#### 3.0 Normal Distribution

#### **3.1 Normal Distribution Manual Process**

If a dataset follows standard normal distribution, then Pr(-1 < Z < 1) determines "the percentage of observation [that] lies within one standard deviation of the mean."

^ mean ( $\mu$ ) + standard deviation ( $\sigma$ ) = 0+1 = 1, mean ( $\mu$ ) - standard deviation ( $\sigma$ ) = 0-1 = -1

- 68% of the data lies within one standard deviation of the mean.
- 95% of the data lies within two standard deviations of the mean.
- 99% of the data lies within three standard deviations of the mean.

#### 3.11 IQR-to-SD ratio

IQR / SD  $\approx 1.34$ 

^ No fixed universally applicable ratio (only testing through one method is not enough to identify if a dataset follows a standard normal distribution or not).

#### 3.12 Quantile-Quantile Normality Plot

Pr(Y < y1) = Pr(X < x1) = Pr(Z < z1) = 1/(n+1)

^ Find the probability of the z-score using 1/(n+1) then use standardization to find the x value.

The x values should be very similar with the y values (thus should have a 45° straight line to be a normal distribution).

#### 3.2 Normal Distribution R Code

# dataset %>% mutate( range1 = if\_else(variablename <= mean(variablename) + sd(variablename) & variablename >= mean(variablename) - sd(variablename),1,0), range2 = if\_else(variablename <= mean(variablename) + 2\*sd(variablename) & variablename >= mean(variablename) - 2\*sd(variablename),1,0), range3 = if\_else(variablename <= mean(variablename) + 3\*sd(variablename) & variablename >= mean(variablename) - 3\*sd(variablename),1,0)) ^ Results will identify the ranges of standard deviations from the mean.

dataset %>% summarize(pct\_range1 = mean(range1), pct\_range2 = mean(range2), pct\_range3 = mean(range3))

^ Finds the number of observations that fall within the mean of the ranges.

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#### 3.2 QQ Normality Plot R Code

dataset <- dataset %>% mutate(rank = rank(variablename), percentile =
rank/(nrow(dataset)+1), stdnorm = qnorm(percentile))
QQ plot (x=stdnorm, y = (variablename - mean(variablename)/sd(variablename)))
^ Allows you to create a QQ plot for any distribution.

ggplot(dataset, aes(sample = variablename)) + geom\_qq() + geom\_qq\_line(colour = "colour", linewidth = #) + labs(x = "Theoretical Returns", y = "Sample Returns") ^ Direct QQ plot for sample.

## 4.0 Sample Probability Distribution4.1 Sample Probability Distribution Manual Process

**Sample Distribution:** Inferring the unknown population mean based on the known sample mean.

**Standard Error:** Standard deviation of the sample (about 10 times smaller than the population standard deviation).

• <u>The more samples we have, the smaller the difference between the sample mean and population mean.</u>

$$\bar{x} \sim \mathcal{N}(\mu_{\bar{x}}, \sigma_{\bar{x}}^2)$$

The mean of the sample means equals to the population mean:

$$\mu_{\overline{x}} = \mu$$

The standard error equals the population's standard deviation divided by the square root of the sample size:

$$\sigma_{\overline{x}} = \frac{\sigma}{\sqrt{n}}$$



**Central Limit Theorem (CLT):** By satisfying the following three conditions, the samples will follow a normal distribution, even if the population does not (z-score can be used):

- We must always be able to draw out multiple samples out of the population.
- We must know the mean and standard deviation of the population,
- The sample size must be sufficiently large (greater than or equal to 30 observations), if it is not a normally distributed population.

#### 4.2 Sample Probability Distribution R Code

```
set.seed(#)
sampleset <- dataset %>% slice_sample(n = #)
^ Seed value can be any value (links the value to a fixed set of sample observations).
```

```
dataset <- data.frame(variablename = numeric())
for (i in 1:10) {
  set.seed(i)
  sample <- dataset %>% slice_sample (n = #)
  temp_var <- mean(sample$variablename)
  mvar <- rbind(mvar, data.frame(avg_var = temp_var, row.names = NULL))}
^ Results allows for estimation of the variable's average variability (taking random samples &
  finding the mean of each, compiling it all together to compare).</pre>
```

## 5.0 Confidence & Significance Levels5.1 Confidence & Significance Levels Manual Testing

<u>Sample Standard Deviation</u>: Where s = sample standard deviation, n = # of observations, and  $\bar{x}$  = sample mean.

$$s = \sqrt{\frac{\sum\limits_{i=1}^{n} (x_i - \bar{x})^2}{n-1}}$$

Point Estimator: Claim of a specific, singular point.

Interval Estimator: Claim of a range that a point may fall into (more confident).



The confidence interval (CI) for an unknown population mean can be written as:

$$\bar{X} - \mathsf{MoE} \leq \mu \leq \bar{X} + \mathsf{MoE}$$

Margin of Error (MoE): Considered a buffer zone.

• MoE = Critical Value \* Standard Error

Standard Error: Population (or sample) Standard Deviation / Sqrt of # of Observations

$$SE = \frac{\sigma (or s)}{\sqrt{n}}$$

Critical Value =  $Z_{\frac{\alpha}{2}}$ 

**Confidence Level:** Measures how confident we are that the calculated interval contains the true (but unobservable) population mean.

- Most Common Confidence Level: 90% / 95% / 99%.
  - $\circ$  e.g., Pr() = 95%
- The greater the confidence level is, the wider the confidence interval is.

#### 5.11 Known SD

Scenario 1:  $\sigma$  is known, confidence level is 95%.

- $\Pr(\bar{x}L < \bar{x} < \bar{x}U) = 95\% (1 \alpha)$
- This means the left over = 5% ( $\alpha$ ), 2.5% on each side ( $\alpha$ /2)  $\leftarrow$  two-tailed.
- That means  $Pr(\bar{x} < \bar{x}L) = 2.5\%$
- Standardization of sample to standard:  $Z = (\bar{x} \mu) / (\sigma / \sqrt{n})$
- Generally: Pr( $-Z \alpha/2 < Z < Z \alpha/2$ )
- Therefore when  $\sigma$  is known:  $\bar{x} Z \alpha/2 * (\sigma / \sqrt{n}) \le \mu \le \bar{x} + Z \alpha/2 * (\sigma / \sqrt{n}) = 1 \alpha$

#### 5.12 Unknown SD

Scenario 2:  $\sigma$  is UNKNOWN, confidence level is still 95%.

• Replace the population sd with the sample sd (s).

$$t = \frac{\bar{X} - \mu}{s/\sqrt{n}}$$



- This will no longer give a standard normal distribution (*Z*). As a result, we use the t-distribution or t-score (t-distribution has fatter tails than the normal distribution).
- By increasing more of the observations, the t-distribution will become more and more like the standard normal distribution.

$$\bar{X} - \frac{t_{\alpha/2}}{\sqrt{n}} \left( \frac{s}{\sqrt{n}} \right) \le \mu \le \bar{X} + \frac{t_{\alpha/2}}{\sqrt{n}} \left( \frac{s}{\sqrt{n}} \right)$$

^ Replaced all of Z with t & all of  $\sigma$  with s.

#### 5.2 T-Statistic R Code

```
mu <- mean(variablename)
sigma <- sd(variablename)
margin_error <- t_value * (sigma / sqrt(sample_size))
ci_lower <- mu - margin_error
ci_upper <- mu + margin_error
^ Finds the MoE and confidence interval of the sample.</pre>
```

#### 6.0 Hypothesis Testing

#### 6.1 Hypothesis Testing Manual Process:

- Null Hypothesis: What we are testing.
- Alternative Hypothesis: Anything but the null hypothesis.
- As words:
  - Null Hypothesis: The mean is equal to \_\_\_\_.
  - Alternative Hypothesis: The alternative hypothesis is not \_\_\_(null hypothesis).
- As symbols:
  - $H_0: \mu =$ \_\_\_\_
  - $H_a: \mu \neq \_$
- NOTE: The null hypothesis can NOT be an inequality.
- Rejection region:
  - $\circ \quad \text{SD known: } Z < \textbf{-} z_{\alpha/2} \text{ or } Z \textbf{>} z_{\alpha/2}$
  - $\circ \quad \text{SD unknown: } t < \textbf{-}t_{\alpha/2} \text{ or } t {>} t_{\alpha/2}$



Two-Tailed Test:

- Alternative hypothesis is <u>non-directional</u>, solely rejecting the null hypothesis.
  - $\circ \quad e.g.,\,H_a:\mu\neq\mu_0$
  - Rejection region is equally split between both tails of the distribution (each area accounts for  $\alpha/2$ ).

One-Tailed Test:

- Alternative hypothesis is <u>directional</u>, in which the entire rejection region ( $\alpha$ ) will be at one tail of the distribution (lower or upper).
  - $\circ \quad e.g., \, H_a: \mu < \mu_0 \; (H_a = \mu > \mu_0)$ 
    - Only one critical value is needed.
    - $Pr(t < t^*) = \alpha$
    - Depending on the H<sub>a</sub>, the p-value probability sign will follow the same way.

P value: Calculating critical value(s) and the area beyond them to determine if  $H_0$  is rejected. R Code: Calculating rejection range using P value, which can be done by:

xbar <- 173.02 mu0 <- 175 s <- 10.95 n <- 162

```
# Calculate t-statistic
t_stat <- (xbar-mu0)/(s/sqrt(n))</pre>
```

```
# Determine p-value
area_lower <- pt(t_stat, df = n-1)
area_upper <- 1- pt(-t_stat, df = n-1)</pre>
```

#return rejection region
area\_lower+area\_upper

Element	One-Sample	Two-Sample
Population Parameter	$\mu$	$\mu_1-\mu_2$
Sample Statistic	$\bar{X}$	$ar{X_1}-ar{X_2}$
Standard Error	$rac{\sigma}{\sqrt{n}}$ or $rac{s}{\sqrt{n}}$	$\sqrt{rac{\sigma_1^2}{n_1}+rac{\sigma_2^2}{n_2}}$ or $\sqrt{rac{s_1^2}{n_1}+rac{s_2^2}{n_2}}$
Degree of Freedom	n-1	$\frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{1}{n_1 - 1}\left(\frac{s_1^2}{n_1}\right)^2 + \frac{1}{n_2 - 1}\left(\frac{s_2^2}{n_2}\right)^2}$